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Cointegration of Economic growth and External balance in Colombia: 1963-2016

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Abstract

In this paper is analyzed the relation between GDP growth and External balance in Colombia for the study period (1963-2016) by using a VECM. Supposing everything else unchanged, we conclude that Colombian external balance granger caused GDP growth and there was indeed a long run relation between both variables. This outcome helps to explain the Colombian GDP growth dynamics over the last fifty years and the impact of trade policy on economic growth.

Keywords

Economic growth; External balance; Trade balance; Colombia; VEC model.

JEL codes: F00, F14, F17, F43, F47.

Resumen

En este artículo se analiza la relación entre crecimiento económico y el saldo de la balanza comercial en Colombia para el periodo de estudio (1963-2016) usando un modelo VEC. Suponiendo todo lo demás constante, se concluye que el saldo de la balanza comercial causo en el sentido de Granger el crecimiento económico colombiano, por lo tanto, existe una relación de equilibrio en el largo plazo entre estas variables. Este resultado contribuye a explicar la dinámica del crecimiento económico colombiano de los últimos cincuenta años y el impacto de la política comercial sobre este.

Palabras claves

Crecimiento económico; Saldo de la balanza comercial; Colombia; Modelo VEC.

JEL: F14, F17, F43.

Resumo

Este artigo analisa a relação entre o crescimento econômico e o saldo da balança comercial na Colômbia para o período de estudo (1963-2016) usando um modelo VEC. Assumindo tudo o mais constante, se conclui que o saldo da balança comercial ocasionado no sentido de Granger o crescimento econômico colombiano, portanto, há uma relação de equilíbrio no longo prazo entre essas variáveis. Esse resultado ajuda a explicar a dinâmica do crescimento econômico colombiano nos últimos cinquenta anos e o impacto da política comercial sobre ele.

Palavras chaves

Crescimento econômico; Saldo da balança comercial; Colômbia; Modelo VEC.

JEL: F14, F17, F43.

1. Introduction

In most of the XX century, Colombian trade policy was aimed to promote economic growth through an increasing External balance and a pretty active intervention on exchange rate and trade markets. That policy was mainly grounded on Prebisch and CEPAL views of Latin American developing countries and their historical low trade terms regarding to developed countries as one of the main causes of their underdevelopment. In Colombia, this trade practices began to change since 1986 in Virgilio Barco's presidential term. However, in 1990, President Gaviria's government started *la Apertura económica* that was a gradual process intended to achieve a greater trade openness and. This policy diminished trade tax in more than 20 per cent, reduce effective protection from 75 per cent to 21 per cent, etc. (Villar et al, 2015). The next four presidents continued this process, but since 2002, Uribe's term and his successor boosted it even more with the sign of more than ten Free trade agreements, including with the USA in 2012 and the EU (Urquijo, 2015). Besides, it is member of regional agreements in trade issues, such as *la Alianza del Pacifico* and *CAN* (Montoya et al, 2016).

To study the relation between GDP growth and External balance, in Colombia for the study period (1963-2016), the last fifty years, it is suitable an economic model for an open economy as Mundell-Fleming (Fleming, 1962). Supposing everything else unchanged, this model sets a positive relation between GDP and a positive External Balance (See equation 1.) The relation between GDP growth and External balance in the year t depends upon value of External balance in year t-1. (See equation 2):

$$Y = \bar{C} + \bar{I} + \bar{G} + EB \text{ (Equation 1)}$$

$$Y = EB$$

$$GDP\ Growth = \frac{(Y_t - Y_{t-1})}{Y_{t-1}} = \frac{(EB_t - EB_{t-1})}{EB_{t-1}}$$

$$\frac{(Y_t - Y_{t-1})}{Y_{t-1}} = GDP\ Growth = \frac{EB_t}{EB_{t-1}} - 1 \text{ (Equation 2)}$$

Based on equation 2, if External balance in year t is zero: $EB_t = 0$, GDP growth is zero. The kind of relation between GDP growth (rate of growth) and External balance (level) in the year t is:

- i. Positive if External balance in year t and External balance in year t-1 are greater than zero or if both are lesser than zero: $(EB_t > 0 \wedge EB_{t-1} > 0) \vee (EB_t < 0 \wedge EB_{t-1} < 0)$
- ii. Negative if External balance in year t is greater than zero and External balance in year t-1 is letter than zero or another way around:

$$(EB_t > 0 \wedge EB_{t-1} < 0) \vee (EB_t < 0 \wedge EB_{t-1} > 0)$$

Since the kind of relation between GDP growth and External balance in year t is not always the same, but also depends on External balance in year t-1 and these latter ones could be actually positive, negative or even zero, it is hard to determine the kind of relation. A VAR model is useful, as in the case of the relation between GDP growth and External balance, it is not easy to distinguish or characterize the structure of relations between two or more variables (Novales, 2014). However, as External balance elasticity of GDP (in level) in year t is unitary, we suppose as our initial hypothesis, that the relation between GDP growth (growth rate) and External balance (in level) was positive in Colombia, in the study period (1963-2016):

$$\varepsilon_{Y,EB_t} = 1 = \frac{\left(\frac{Y_t - Y_{t-1}}{Y_{t-1}}\right)}{\left(\frac{EB_t - EB_{t-1}}{EB_{t-1}}\right)}$$

In fact, ε_{Y,EB_t} of Colombia was positive: 0.15, but relatively inelastic, on average, in the study period.

2. Econometric analysis.

2.1. VAR Model

An econometric analysis is done, intended to, clarify a bit more the relation between External balance and GDP growth. This is done through a special kind of VAR model (Vector auto regression) called VEC (Vector error correction) both make it possible to find the multiple linear relations among time series, but besides the latter provides information about long-run equilibrium between variables and how fast they come back to it, after short-run shocks (Novales, 2014).

Stock and Watson (2001) warn us, one of the main problems of VAR models is its interpretation, since identification problem normally arises; it is difficult to discern between correlation and causation. Solving that problem is not just a statistical issue, but it requires adding a theoretical framework, that in this work was already presented above.

According to EViews (2018), for the case of a two-time series system, one cointegrating equation between them ($y_{2t} = \aleph_1 y_{1t} + c_1$) and one lagged difference, the first equation of VEC model is:

$$\Delta y_{1,t} = \alpha(y_{1,t-1} - \aleph_{11}y_{2,t-1} - c_1) + \aleph_{12}\Delta y_{1,t-1} + \aleph_{13}\Delta y_{2,t-1} + K_1 + \varepsilon_{1t}$$

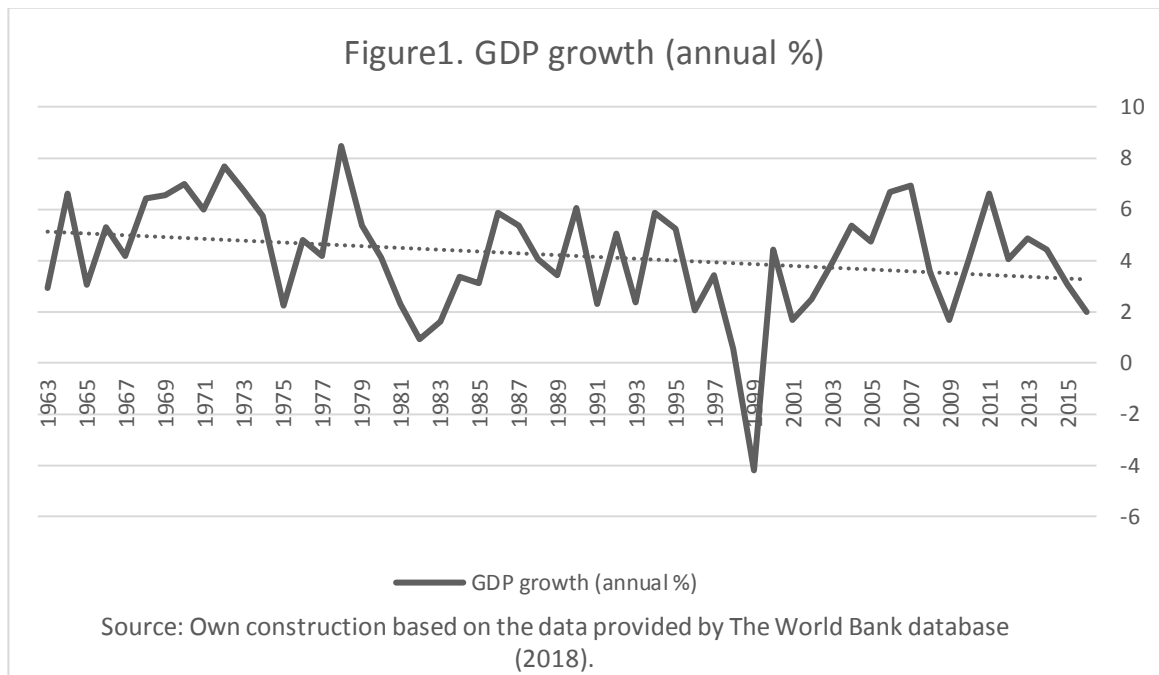
(Equation 3)

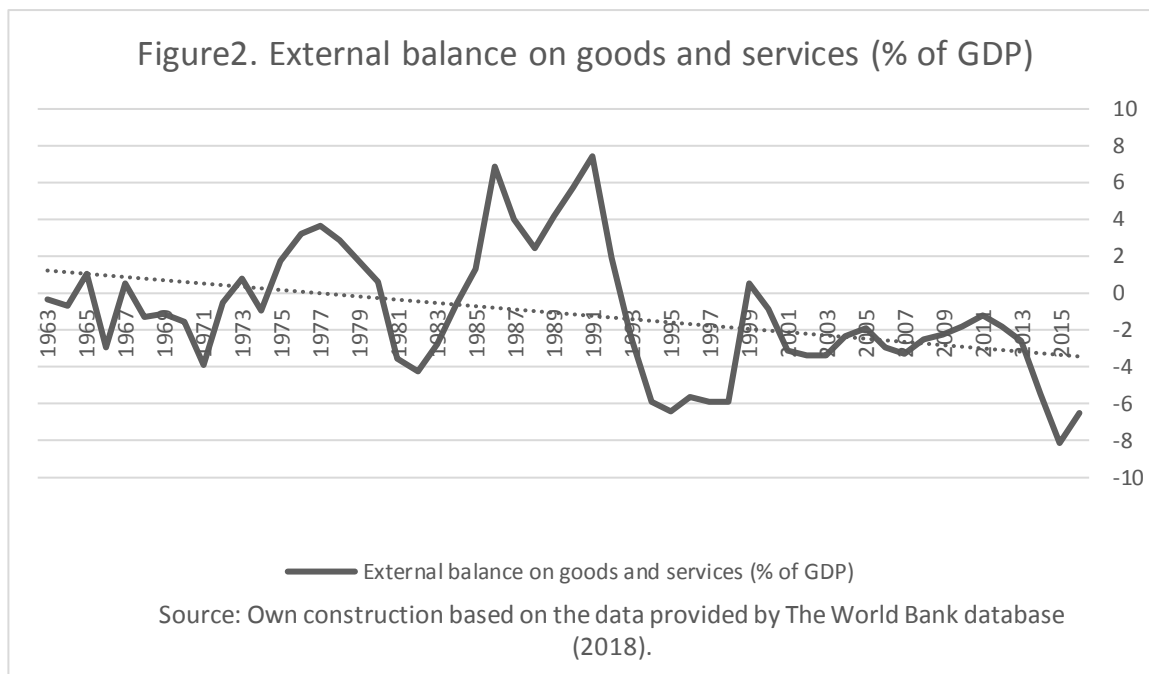
where, $\Delta y_{1,t-1}$ is the first difference of the first endogenous variable; $y_{1,t-1}$ is the first endogenous variable with one lag; \aleph_{11} is coefficient of the sole cointegration equation; c_1 is constant term of the sole cointegration equation; \aleph_{12} is coefficient of coefficient error term of the first variable with one lag; \aleph_{13} is coefficient of coefficient error term of the second variable with one lag; ε_{1t} are residues of the first equation of VEC model; K_1 is constant term of the first equation of VEC model ; α is coefficient of the error correction term in long run equilibrium of the first variable, it must be <0 and measures speed of adjustment of the first variable towards long-run equilibrium (EViews, 2018).

The second equation is analogous to the first one. Both series must satisfy the next conditions.

1. y_{1t} and y_{2t} are cointegrating processes of order 1: $I(1)$:
2. ε_{1t} and ε_{2t} are white noise processes, which means that over time their expected values are constant, in this case equal to zero: $E[\varepsilon_{1t}] = 0$ and $E[\varepsilon_{2t}] = 0$; but also their variances are constant across time: $\sigma_{\varepsilon_{1t}}^2 = c$ and $\sigma_{\varepsilon_{2t}}^2 = c$, where $c \in R^+$, including zero.
3. $\text{Cov}(\varepsilon_{1t}, \varepsilon_{2t}) = \sigma_{\varepsilon_{1t}\varepsilon_{2t}}$

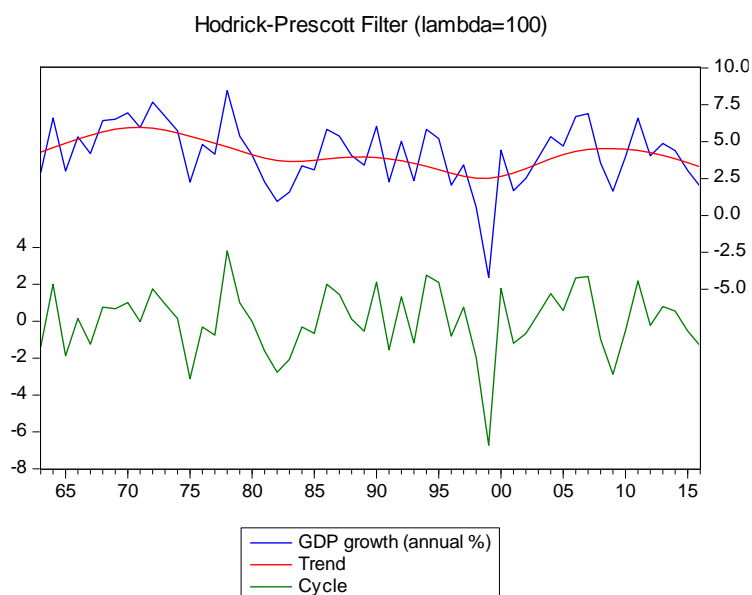
2.2 Graphical analysis





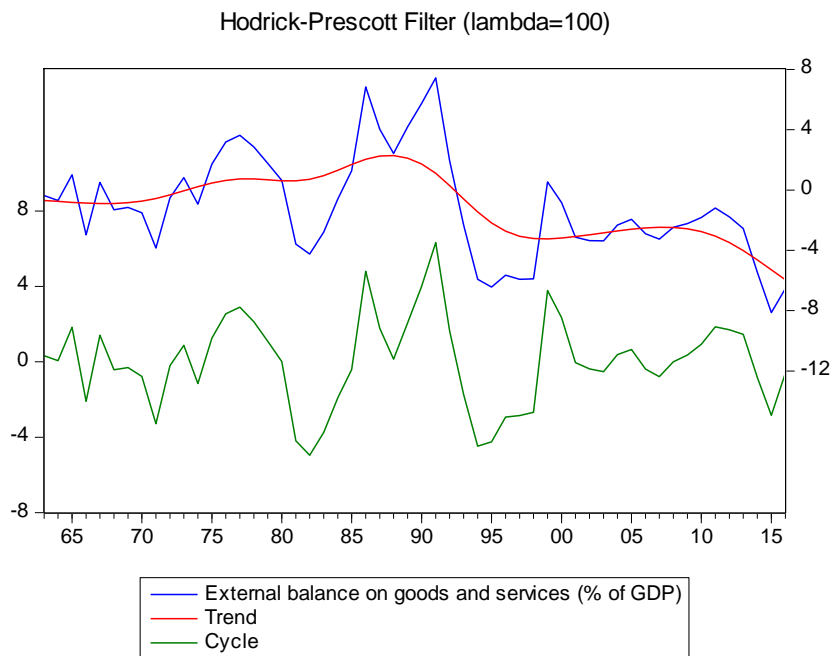
Based on Figure 1 and Figure 2, both series seem to have negative trends. These ones could be deterministic or stochastic. Applying Hodrick-Prescott filter (with a power of 2 and a lambda equals to 100) to both series to decompose and better observe their trends and cycle components, produces Figure 3 y Figure 4:

Figure 3.



Source: World Bank (2018). Estimation made on EViews.

Figure 4.



Source: World Bank (2018). Estimation made on EViews.

In this work, it is assumed that both trends are stochastic. In fact, most economic series also are it, thus it is a weaker assumption than deterministic trend one. Besides, for any amount of non-infinite data there is a deterministic trend and stochastic trend that fits the data equally well (Hamilton, 1994). A Stochastic trend mean involves non-stationarity (Hamilton, 1994).

2.3 Unit root test

According to Hamilton (1994), the ADF test (augmented Dickey-Fuller) contrasts the null hypothesis that the series has a unit root, which means that it is non-stationary. The alternative hypothesis asserts that the series does not have a unit root, consequently the series is stationary. Only in case of rejection of the null hypothesis at a level of significance: α ; it is possible to assert that there is significant statistical evidence to assert that the series is stationary and if so does, it is possible to use the VAR model,

avoiding doing a spurious regression. In this work, the chosen significance level is: $\alpha = 0.05$.

In Figure 3, there is a clear tendency; therefore, we test stationery of this series by using Augmented Dickey Fuller (ADF) test including a trend and an intercept in the unit-root test. Specifying the test for a max of three lags to the GDP growth series, we get a p-value = 0.0712. Hence, at a significance level: $\alpha = 0.05$, it can be said that there is significant statistically evidence, to state that the GDP growth series is non-stationary (See Annex 1).

In Figure 4, we test formally the stationery of this series by using Augmented Dickey Fuller (ADF) test including a trend and an intercept in the unit-root test. Specifying the test for a max of three lags to the External balance series, we get a p-value = 0.058. Hence, at a significance level: $\alpha = 0.05$, it can be said that there is significant statistically evidence, to state that the External balance series is non-stationary (See Annex 2).

2.4 Integration order

A non-stationary series (y_{it}) is integrating of order d: I (d) if it is necessary take d differences to make it stationary. In our case, if we take the first difference of each series: the GDP growth and the External balance series and apply ADF test with trend and intercept (three lags) at the 0.05 level, we get two new stationary series (See Annex 3), thus both the GDP growth and the External balance series are integrating of order 1: I (1). A VEC model with k endogenous variables should have k-1 cointegrating relations (EViews, 2018).

2.5 Akaike information criterion

In order to select the optimal number of lags for our model VEC, it is suitable to use the Akaike information criterion (AIC), which estimates how much information a stochastic process provides in comparison to other ones (Novales, 2014), To get this criterion, it is necessary to run a VAR model, in spite of the fact that both series are non-stationary. The optimal number of lags in VEC models is one lesser than that VAR

model. The AIC results that the most suitable VAR or that one which provides the most amount of information is a VAR model with two lags: VAR (2) (See Annex 4), therefore our VEC model should have an optimal number of lags equal to one: VEC (1).

2.6 The cointegration test of Johansen

The cointegration test of Johansen estimates whether a pair or more series are cointegrated one another or not. This test can have many null hypotheses; each one asserts that does not exist a determined cointegration level. In our case, given the fact that the both series have a stochastic trend, it must be tested by assuming that there is deterministic trend in data and the sole cointegrating equation has intercept and VAR test (EViews, 2018). The results of this tests does not reject the null hypotheses at the 0.05 level, consequently the model is cointegrated (See Annex 5.), which is ideal because it guarantees that the cointegration rank of differentiated series is not zero (Enders, 2003).

2.7 Estimated VAR model

After having checked that the series are non-stationary, have the same integrating order: I (1) and the optimal number of lags of VEC model, it is possible to run a VEC (1) model. Given the fact that both series have stochastic trend, the VEC model must be estimated without trend and with intercept in both Cointegrating equation and VAR (EViews, 2018) (See Annex 4.). The estimated VEC (1) model is:

$$\Delta GDP\ growth_{1,t} = 0.064(External\ balance_{t-1} - 8.07GDP\ growth_{t-1} - 35.35) + 0.24\Delta External\ balance_{,t-1} - 0.14\Delta GDP\ growth_{t-1} - 0.052 + \varepsilon_{1t}$$

(Equation 4)

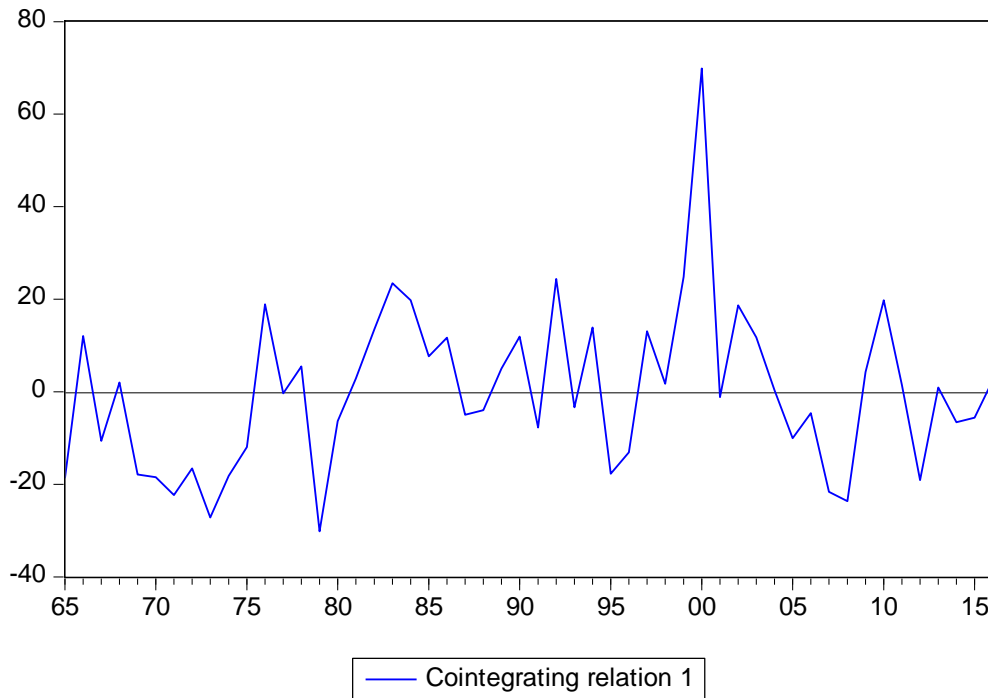
$$\Delta External\ balance_{2,t} = 0.002(External\ balance_{t-1} - 8.07GDP\ growth_{t-1} + 35.35) + 0.059\Delta External\ balance_{,t-1} + 0.051\Delta GDP\ growth_{t-1} - 0.10 + \varepsilon_{2t}$$

(Equation 5)

The sole estimated cointegrating equation is depicted in Figure 5.:

$$\overline{External\ balance_{t-1t-1}} = 8.07GDP\ growth_{t-1} + 35.35 \quad \textbf{(Equation 6)}$$

Figure 5.



Source: World Bank (2018). Estimation made on EViews.

3. Assumptions

It is broadly known that in VAR models, parameters and R^2 coefficient are not usually interpreted (Sosa, 2016). It must be checked that the estimated model is not a spurious regression, that is to say the series reflect a significant implicit relation between GDP

growth and External balance, thus, it is necessary the accomplishment of some assumptions of VEC model and their statistical significance.

3.1. Correlograms

First, we see the correlograms of residuals of the equations of our model until 12 lags, in order to check that there is not autocorrelation among them (See Annex 6.). We find that in an equation appears to be autocorrelation of thirteen order, however other Autocorrelation tests discard this possibility. The VEC Portmanteau test for autocorrelations have as null hypothesis that there is not serial correlation for each lag. To verify the non-autocorrelation of residuals assumption, the VEC estimated model is tested, in our case up to twenty lags, resulting that the model fulfills this assumption in the Portmanteau tests, at a significance level: $\alpha = 0.05$, (See Annex 7.). For the rest of lags, there is not any problem, thus this assumption is completely fulfilled.

3.2. Stability of VEC model

In order to verify the stability of the VEC model across the time, through the inverse roots of the characteristic AR polynomial test, it is vital to know that in a VEC model only $k-r$ roots should be equal to unity, where k is the number of endogenous variables and r is the number of cointegrating relations (EViews, 2018). In the estimated VEC model $k-r$ is equal to one. Testing roots of characteristic polynomial of the estimated VECM, at a significance level: $\alpha: 0.05$, this assumption is fulfilled, since only one root is equal to one (See Annex 8.).

An assumption of VEC model is that residuals are distributed normally. To verify it, we use VEC Residual Normality Tests (according to Orthogonalization: Cholesky (Lutkepohl) criterion), validating also the fulfillment of this assumption at a significance level: $\alpha = 0.05$, (See Annex 9.). Regarding the assumption of homokedasticity of residuals, to verify it, it is used the white heteroskedasticity test with no cross terms and the white heteroskedasticity test with cross terms at a significance level: $\alpha = 0.05$, the estimated VEC model also fulfills assumption of homokedasticity of residuals (See Annex 10.)

4. Granger causality

In a spurious model the variables are inferred statistically as if they were causally correlated, but they are actually independent to each other. That kind of model is not desirable at all, in order to verify if it happens or not, it is necessary to perform the Granger Causality Test (Novales, 2014).

A variable X_t causes in the sense of Granger to another Y_t , if, past values of X_t : (X_{t-1} , X_{t-2} and so forth) if X_t is able to predict future values of Y_t : (Y_{t+1} , Y_{t+2} and so on). The Granger test tells us, if that predictive capacity of the past values of X_t is statistically significant or not, at a level of significance α (Novales, 2014). The null hypothesis of the Granger test says: X_t does not Granger cause Y_t . The alternative hypothesis of the Granger test says: X_t does Granger cause Y_t (Lin, 2008). Therefore, running this test in the estimated VEC model will help to identify whether External balance granger caused the distribution of the GDP growth (See Table 1.):

Table 1. VEC Granger causality test.

VEC Granger Causality/Block Exogeneity Wald Tests
Date: 04/01/18 Time: 22:05
Sample: 1963 2016
Included observations: 52

Dependent variable: D(EXTERNAL_BALANCE_ON_GOOD)

Excluded	Chi-sq	df	Prob.
D(GDP_GRO...	0.099620	1	0.7523
All	0.099620	1	0.7523

Dependent variable: D(GDP_GROWTH__ANNUAL__)

Excluded	Chi-sq	df	Prob.
D(EXTERNAL...	3.997039	1	0.0456
All	3.997039	1	0.0456

Data Source: World Bank (2018). Estimation made on EViews.

On the one hand, as the p-value (0.7523) in the first regression is greater than the level of significance (0.05), it is possible to say that the first difference (Variations of year $t-1$ to t) of GDP growth did not Granger cause the first difference of External balance, in the period 1963-2016. On the other hand, as the p-value (0.0456) of the second regression is less than the level of significance (0.05), it is possible to say that the first difference of External balance did Granger cause the first difference of GDP growth, in the period 2000-2014.

The interpretation of the results of the econometric model, especially Granger's causality test, is the following:

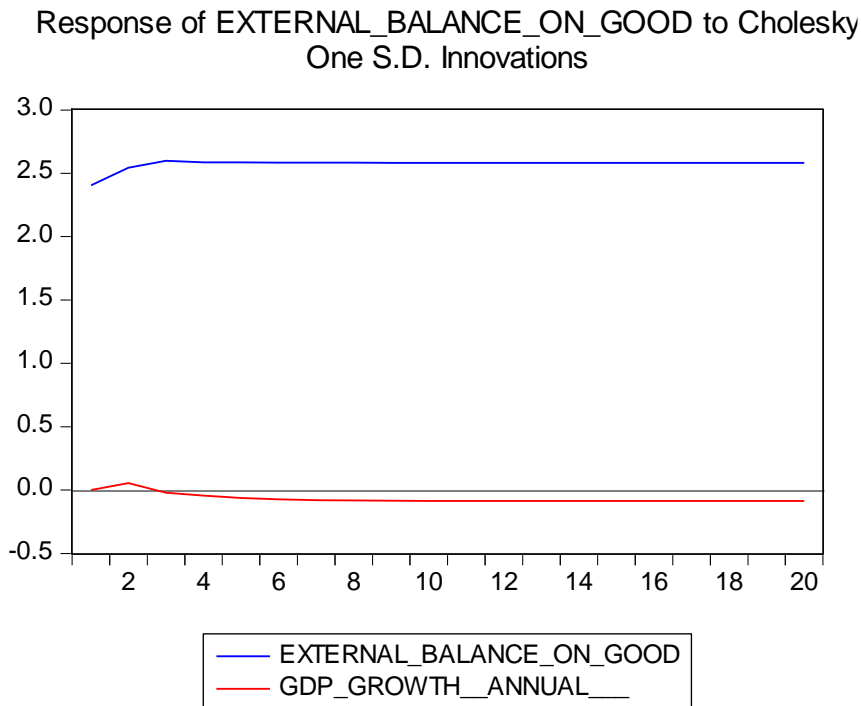
Variations in GDP growth did not explain in a temporary sense the External balance, in the study period.

Variations in the External balance did explain in a temporary sense the variations in the GDP growth, in other words, they explain, at least partially (in a temporary sense), the GDP growth.

5. Impulse response functions.

The residuals of the VEC model: ε_{1t} and ε_{2t} can be thought as deviations (impulses, random shocks or innovations) of each one of the endogenous variables from a perfect estimation (Stock & Watson, 2001). The impulse response functions of the model express how the behavior of one endogenous variable changes against a simulated random shock (innovation) of the other variables' residuals or theirs. It is supposed that all other errors are equal to zero and that a particular error of the estimated VECM (random shock) returns to zero afterwards, in the long-run (Stock & Watson, 2001). In this case, it is supposed responses up to 20 periods after the random shock and the decomposition method used is Cholesky (DOF adjusted), with innovations equal to one standard deviation:

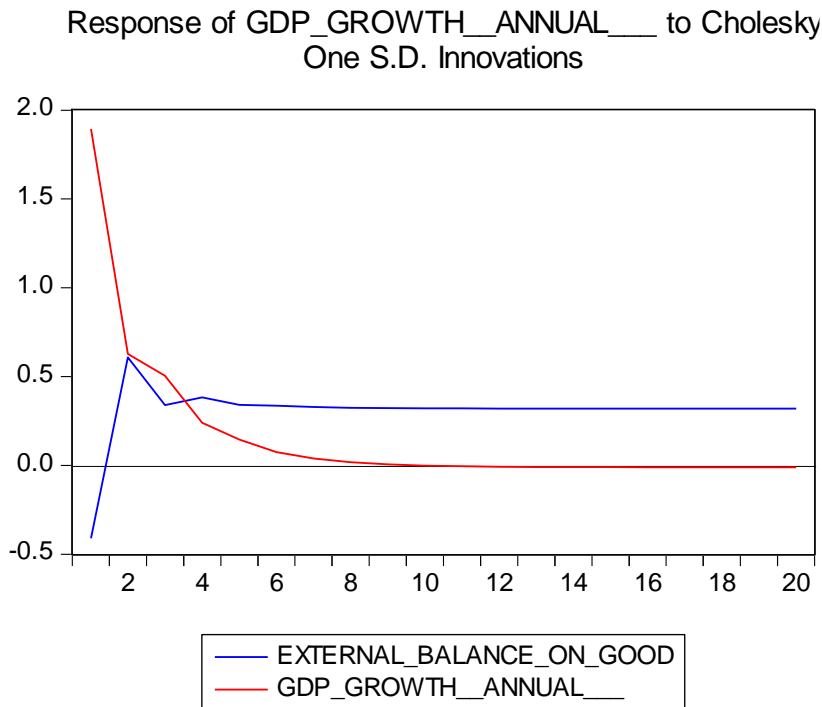
Figure 6. Impulse response functions.



Source: World Bank (2018). Estimation made on EViews.

The Figure 6 shows the effect of an unexpected one standard deviation increase in the External balance (blue line) and the GDP growth (red line) on the External balance. The Figure 6 depicts as a shock or increase of External balance in one standard deviation increases the External balance, having a permanent effect on itself through the 20 periods, that is to say, the long-run. As for, the GDP growth, their innovations have a slight positive effect on External balance initially, but from third lag this effect changes and becomes a negative one, but its effect is too weak though. A random shock on External balance lasts up to twentieth-post-impulse period. In other words, innovations spread in the long run.

Figure 7. Impulse response functions.



Source: World Bank (2018). Estimation made on EViews.

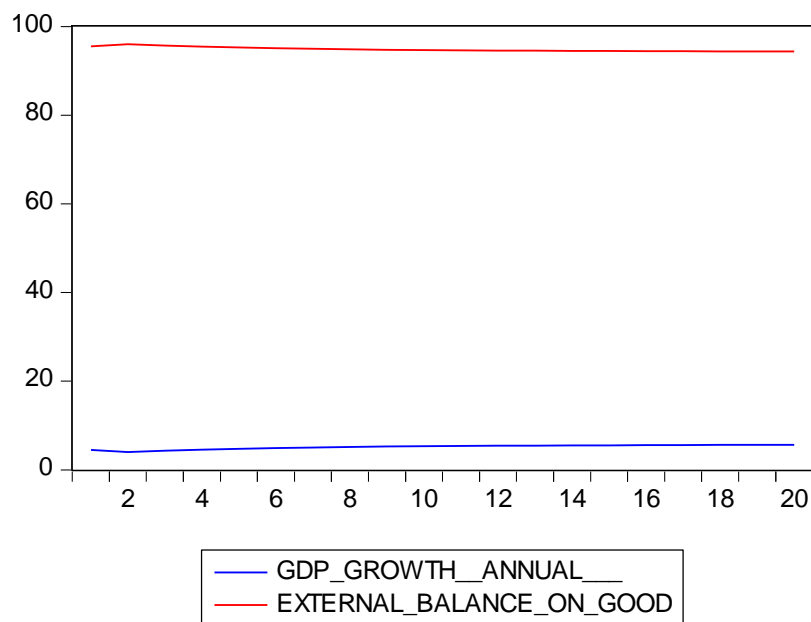
The Figure 7 shows the effect of an unexpected one standard deviation increase in The GDP growth (red line) and the External balance (blue line) on the GDP growth. As for, GDP growth, as expected its own innovations have a strong positive effect on GDP growth itself initially, which decrease abruptly from the second lag but keeps on being positive up to the seventh period, from where begins to wane and becomes almost nil, in other words has an effect in medium-run. On the other hand, initially, an innovation of the External balance increases mildly GDP growth, but having a permanent effect through the 20 periods though, that is to say, a random shock of External balance lasts up to twentieth-post-impulse period. Put simply, innovations of the External balance on the GDP growth spread in the long run.

6. Variance Decomposition

According to Novales (2014), Variance decomposition helps to provide an estimation of size of prediction error of each variable, in terms of each of other variables. Errors inevitably increase with prediction horizon. It is, therefore, a way of making inference about the intertemporal relations between the variables that compose VEC model. The components of each variance are expressed in percentage terms (See Figure 8 and Figure 9.).

Figure 8

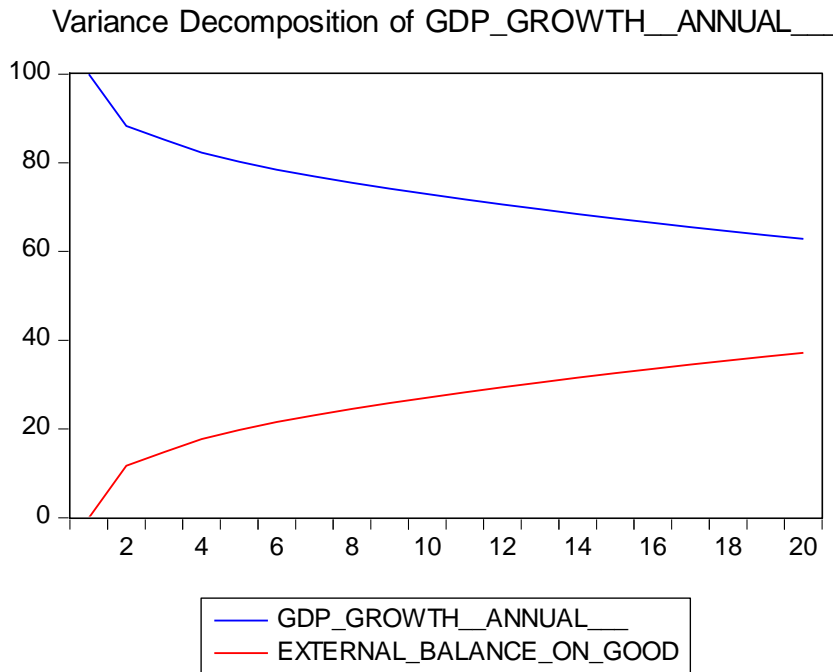
Variance Decomposition of EXTERNAL_BALANCE_ON_GOOD



Source: World Bank (2018). Estimation made on EViews.

The estimated variance for the variable the External balance (See Figure 8.) is almost explained entirely (95 per cent) by their own shock over all the lag period. The other percentage (roughly 5 per cent) of variance depends on the GDP growth

Figure 9.



Source: World Bank (2018). Estimation made on EViews.

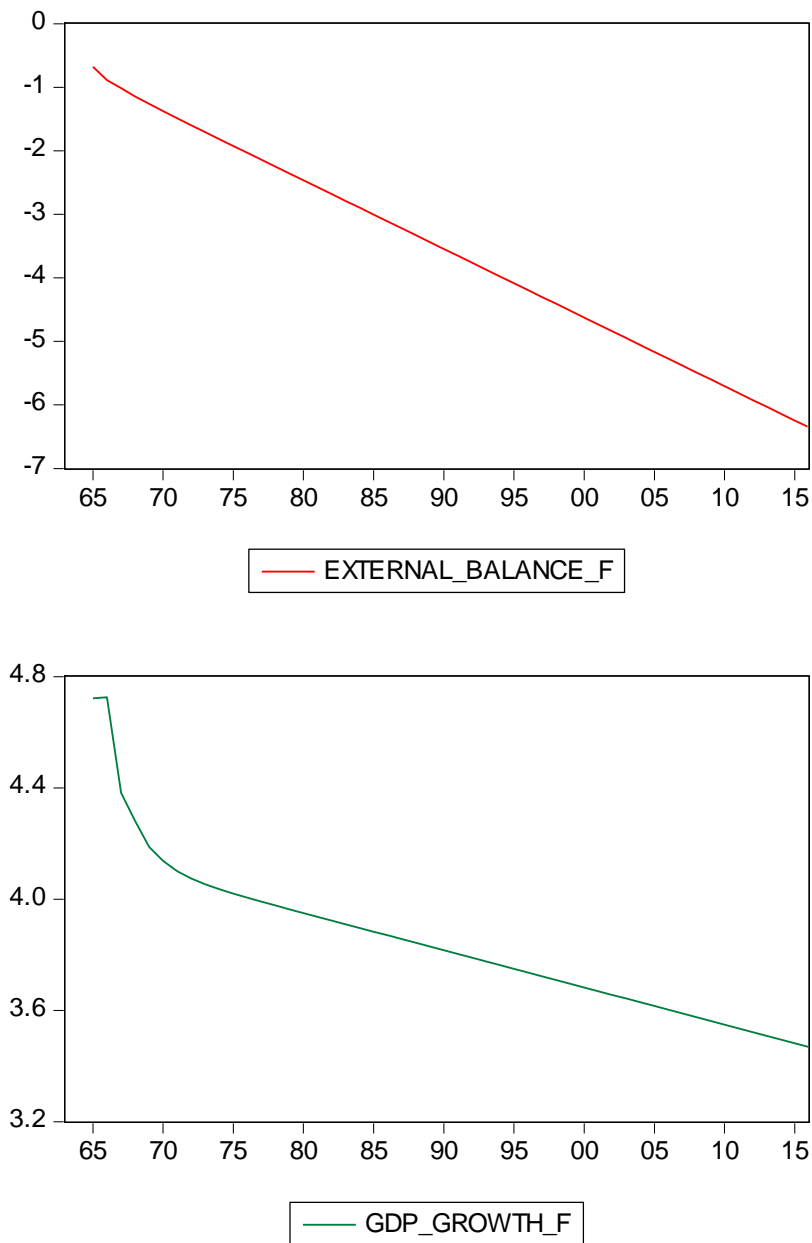
As far as the GDP growth is concerned, the contribution of variables, to the variance of the GDP growth changes over time (See Figure 8.), so that, after the fifth lag, about 20 per cent of its variance is explained by the External balance. This percentage skyrocketed up to 37 per cent in the twentieth period. On average, the External balance explains 30 per cent of the variance of the GDP growth. Up to the other 70 per cent, on average, of its own variance is explained by the variable itself: the GDP growth.

7. Forecasting

Applying a dynamic forecast to the estimated model for the sample period: 1963-2016 and by including actual values for the sample observations and the coefficient of uncertainty of standard errors (to reduce the largest source of forecast error, as the

forecast residuals are not known), a set of forecast values are made for both of the series:

Figure 10. Forecasting of GDP growth and External balance.



Source: World Bank. Estimation made on EViews.

Some indicators such as: RMSE, MAE, MAPE and Theil index; which are pretty helpful to evaluate the suitability of the estimated model. A simple guideline to grasp

the meaning of these measures is the lower values are, the better forecasts the model produces (Woschnagg & Cipan, 2004):

Table 2. Forecasting evaluation.

Forecast Evaluation
Date: 04/02/18 Time: 04:00
Sample: 1963 2016
Included observations: 54

Variable	Inc. obs.	RMSE	MAE	MAPE	Theil
D(GDP_GROWT...	54	4.589709	3.722005	115.0524	0.715251
D(EXTERNAL_B...	54	4.588735	4.126745	107.6534	0.743162

RMSE: Root Mean Square Error
MAE: Mean Absolute Error
MAPE: Mean Absolute Percentage Error
Theil: Theil inequality coefficient

Source: World Bank. Estimation made on EViews.

Theil index is a good benchmark to help gauge results. The closer Theil inequality coefficients are to 1, the lesser ability to forecast the model has. On the other hand, the closer Theil inequality coefficients are to 0, the greater ability to forecast the model has (Woschnagg & Cipan, 2004). In the estimated VEC model case (See Table 2.), the Theil index of The GDP growth is 0.71, (Equation (4)), that means that fifth equation's ability is far better forecast than the estimated equation for the External balance (Equation 5), whose Theil index is a bit higher 0.74.

The impulse response functions, the variance decomposition analysis and the forecasting process suggest that External balance had a role in the dynamics of the GDP growth in the long-run. In the estimated VEC model, External balance's random shocks have positive impacts in increasing GDP growth of the first quintile, but of mild intensity though.

As has been said at the beginning of this paper, GDP growth depends not only on External balance, but also on other aggregate variables such as consume, investment and public expenditure. They have not been taken them into account in the econometric model, mainly because the purpose of this work is to analyze only the relation between

GDP growth and External balance. Furthermore, by not taking into account these variables, the model VECM avoids losing degrees of freedom ($n-k$; k is the number of parameters) to the most. By doing so, we make the estimated model as parsimonious as possible.

8. Discussion and conclusions:

Now, we are able to answer the fundamental question, with which this work started: What was the relation between GDP growth and External balance in Colombia, in the long run, during the period: 1963-2016? The econometric analysis helped to determine the validity of our initial hypothesis, in the period 1963-2016, the GDP growth, and the External balance had a positive relation in the long-run. After that, it was checked out that the External balance did Granger cause the GDP growth, this entails that during the study period, there was a direct and a causal relation (in a temporary sense) between External balance and GDP growth, but its effect was mild, though.

Given the fact, that the definition of External balance, according to the World Bank is difference between Exports of goods and services minus Imports of goods and services (World Bank, 2018). A change of sign of the average of the External balance involves an external deficit on average, and therefore as the External Balance Granger causes GDP growth, an average external deficit must have a negative impact on the economic growth. What joint with the external balance Elasticity of GDP, the estimated impulse response analysis and the granger casualty test, support the validity of our initial hypothesis,

A positive External balance or a trade surplus could boost economic growth, through several ways, for example an increase of exports, better terms of trade, a decrease of imports, or a rise of productivity of factors in export sectors that rises their income shares, (Casas & Guzman, 2015). Actually, External balance depends, at least partially, on foreign GDP, through demand of exports, but also depends on local GDP, negatively at least partially, through demand of imports (Oliveros & Huertas, 2015) and in turn

local GDP depends positively of imports, since a percentage of imports is used as input in local production, etc. That's why having chosen a VAR model was a suitable option, since it deals with both External balance and GDP growth as endogenous variables, helping us to deal with problems of endogeneity, at least empirically. In the study period, in Colombia, External balance caused in a temporary sense to GDP growth, and not another way around.

9. References.

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Annex 1.

Null Hypothesis: GDP_GROWTH__ANNUAL___ has a unit root

Exogenous: Constant, Linear Trend

Lag Length: 3 (Fixed)

	t-Statistic	Prob.*
Augmented Dickey-Fuller test statistic	-3.342354	0.0712
Test critical values:		
1% level	-4.152511	
5% level	-3.502373	
10% level	-3.180699	

*MacKinnon (1996) one-sided p-values.

Augmented Dickey-Fuller Test Equation

Dependent Variable: D(GDP_GROWTH__ANNUAL___)

Method: Least Squares

Date: 04/01/18 Time: 13:44

Sample (adjusted): 1967 2016

Included observations: 50 after adjustments

Variable	Coefficient	Std. Error	t-Statistic	Prob.
GDP_GROWTH__ANNUAL___(-1)	-0.723303	0.216405	-3.342354	0.0017
D(GDP_GROWTH__ANNUAL___(-1...	0.017742	0.200614	0.088439	0.9299
D(GDP_GROWTH__ANNUAL___(-2...	0.168121	0.183881	0.914291	0.3655
D(GDP_GROWTH__ANNUAL___(-3...	0.111574	0.146239	0.762955	0.4496
C	3.895373	1.323196	2.943913	0.0052
@TREND("1963")	-0.031587	0.022083	-1.430361	0.1597
R-squared	0.355368	Mean dependent var	-0.067246	
Adjusted R-squared	0.282115	S.D. dependent var	2.470879	
S.E. of regression	2.093529	Akaike info criterion	4.427746	
Sum squared resid	192.8460	Schwarz criterion	4.657189	
Log likelihood	-104.6936	Hannan-Quinn criter.	4.515119	
F-statistic	4.851209	Durbin-Watson stat	2.015095	
Prob(F-statistic)	0.001280			

Annex 2.

Null Hypothesis: EXTERNAL_BALANCE_ON_GOOD has a unit root

Exogenous: Constant, Linear Trend

Lag Length: 3 (Fixed)

	t-Statistic	Prob.*
Augmented Dickey-Fuller test statistic	-3.432895	0.0584
Test critical values:		
1% level	-4.152511	
5% level	-3.502373	
10% level	-3.180699	

*MacKinnon (1996) one-sided p-values.

Augmented Dickey-Fuller Test Equation

Dependent Variable: D(EXTERNAL_BALANCE_ON_GOOD)

Method: Least Squares

Date: 04/01/18 Time: 13:46

Sample (adjusted): 1967 2016

Included observations: 50 after adjustments

Variable	Coefficient	Std. Error	t-Statistic	Prob.
EXTERNAL_BALANCE_ON_GOOD(-1)	-0.429597	0.125141	-3.432895	0.0013
D(EXTERNAL_BALANCE_ON_GOOD(-1...	0.309175	0.145867	2.119559	0.0397
D(EXTERNAL_BALANCE_ON_GOOD(-2...	0.150008	0.145772	1.029053	0.3091
D(EXTERNAL_BALANCE_ON_GOOD(-3...	0.139609	0.145994	0.956266	0.3442
C	0.978733	0.686201	1.426306	0.1608
@TREND("1963")	-0.050029	0.023005	-2.174648	0.0351
R-squared	0.235292	Mean dependent var	-0.071439	
Adjusted R-squared	0.148393	S.D. dependent var	2.307950	
S.E. of regression	2.129835	Akaike info criterion	4.462132	
Sum squared resid	199.5926	Schwarz criterion	4.691575	
Log likelihood	-105.5533	Hannan-Quinn criter.	4.549506	
F-statistic	2.707661	Durbin-Watson stat	1.846787	
Prob(F-statistic)	0.032253			

Annex 3.

Null Hypothesis: D(GDP_GROWTH__ANNUAL____) has a unit root
 Exogenous: Constant, Linear Trend
 Lag Length: 3 (Fixed)

	t-Statistic	Prob.*
Augmented Dickey-Fuller test statistic	-4.395403	0.0052
Test critical values:		
1% level	-4.156734	
5% level	-3.504330	
10% level	-3.181826	

*MacKinnon (1996) one-sided p-values.

Augmented Dickey-Fuller Test Equation
 Dependent Variable: D(GDP_GROWTH__ANNUAL____,2)
 Method: Least Squares
 Date: 04/01/18 Time: 16:20
 Sample (adjusted): 1968 2016
 Included observations: 49 after adjustments

Variable	Coefficient	Std. Error	t-Statistic	Prob.
D(GDP_GROWTH__ANNUAL____(-1))	-1.962626	0.446518	-4.395403	0.0001
D(GDP_GROWTH__ANNUAL____(-1),...	0.473417	0.372826	1.269806	0.2110
D(GDP_GROWTH__ANNUAL____(-2),...	0.253027	0.273047	0.926679	0.3593
D(GDP_GROWTH__ANNUAL____(-3),...	0.098387	0.148521	0.662442	0.5112
C	0.093470	0.769448	0.121477	0.9039
@TREND("1963")	-0.005469	0.023846	-0.229364	0.8197
R-squared	0.713506	Mean dependent var		0.000812
Adjusted R-squared	0.680192	S.D. dependent var		4.170615
S.E. of regression	2.358547	Akaike info criterion		4.668247
Sum squared resid	239.1980	Schwarz criterion		4.899898
Log likelihood	-108.3720	Hannan-Quinn criter.		4.756135
F-statistic	21.41804	Durbin-Watson stat		2.029380
Prob(F-statistic)	0.000000			

Null Hypothesis: D(EXTERNAL_BALANCE_ON_GOOD) has a unit root
Exogenous: Constant, Linear Trend
Lag Length: 3 (Fixed)

	t-Statistic	Prob.*
Augmented Dickey-Fuller test statistic	-4.284780	0.0071
Test critical values:		
1% level	-4.156734	
5% level	-3.504330	
10% level	-3.181826	

*MacKinnon (1996) one-sided p-values.

Augmented Dickey-Fuller Test Equation
Dependent Variable: D(EXTERNAL_BALANCE_ON_GOOD,2)
Method: Least Squares
Date: 04/01/18 Time: 16:21
Sample (adjusted): 1968 2016
Included observations: 49 after adjustments

Variable	Coefficient	Std. Error	t-Statistic	Prob.
D(EXTERNAL_BALANCE_ON_GOOD(-1))	-1.224640	0.285812	-4.284780	0.0001
D(EXTERNAL_BALANCE_ON_GOOD(-1),...	0.357315	0.243285	1.468709	0.1492
D(EXTERNAL_BALANCE_ON_GOOD(-2),...	0.246304	0.198485	1.240918	0.2214
D(EXTERNAL_BALANCE_ON_GOOD(-3),...	0.189125	0.143424	1.318641	0.1943
C	0.077278	0.756744	0.102119	0.9191
@TREND("1963")	-0.007913	0.023574	-0.335659	0.7388
R-squared	0.469380	Mean dependent var	-0.038569	
Adjusted R-squared	0.407680	S.D. dependent var	2.998564	
S.E. of regression	2.307764	Akaike info criterion	4.624713	
Sum squared resid	229.0083	Schwarz criterion	4.856364	
Log likelihood	-107.3055	Hannan-Quinn criter.	4.712601	
F-statistic	7.607464	Durbin-Watson stat	1.856849	
Prob(F-statistic)	0.000034			

Annex 4.

VAR Lag Order Selection Criteria

Endogenous variables: EXTERNAL_BALANCE_ON_GOOD GDP_GROWTH__ANNUAL__

Exogenous variables: C

Date: 04/01/18 Time: 14:39

Sample: 1963 2016

Included observations: 49

Lag	LogL	LR	FPE	AIC	SC	HQ
0	-239.4846	NA	65.41621	9.856513	9.933730	9.885809
1	-209.1328	56.98691*	22.32004	8.780932	9.012584*	8.868821*
2	-204.5973	8.145554	21.86184*	8.759071*	9.145157	8.905552
3	-203.5856	1.734236	24.75994	8.881045	9.421565	9.086118
4	-203.0629	0.853407	28.66436	9.022976	9.717930	9.286640
5	-201.9250	1.764882	32.44674	9.139797	9.989185	9.462053

* indicates lag order selected by the criterion

LR: sequential modified LR test statistic (each test at 5% level)

FPE: Final prediction error

AIC: Akaike information criterion

SC: Schwarz information criterion

HQ: Hannan-Quinn information criterion

Annex 5

Date: 04/01/18 Time: 16:28

Sample (adjusted): 1965 2016

Included observations: 52 after adjustments

Trend assumption: Linear deterministic trend

Series: EXTERNAL_BALANCE_ON_GOOD GDP_GROWTH__ANNUAL__

Lags interval (in first differences): 1 to 1

Unrestricted Cointegration Rank Test (Trace)

Hypothesized No. of CE(s)	Eigenvalue	Trace Statistic	0.05 Critical Value	Prob.**
None *	0.192052	18.58335	15.49471	0.0166
At most 1 *	0.134212	7.493971	3.841466	0.0062

Trace test indicates 2 cointegrating eqn(s) at the 0.05 level

* denotes rejection of the hypothesis at the 0.05 level

**MacKinnon-Haug-Michelis (1999) p-values

Unrestricted Cointegration Rank Test (Maximum Eigenvalue)

Hypothesized No. of CE(s)	Eigenvalue	Max-Eigen Statistic	0.05 Critical Value	Prob.**
None	0.192052	11.08938	14.26460	0.1498
At most 1 *	0.134212	7.493971	3.841466	0.0062

Max-eigenvalue test indicates no cointegration at the 0.05 level

* denotes rejection of the hypothesis at the 0.05 level

**MacKinnon-Haug-Michelis (1999) p-values

Unrestricted Cointegrating Coefficients (normalized by b*S11*b=I):

EXTERNAL_B...	GDP_GROWTH__ANNUAL__
-0.073307	0.592173
0.324604	-0.011917

Unrestricted Adjustment Coefficients (alpha):

D(EXTERNAL_...	-0.037822	-0.846336
D(GDP_GRO...	-0.880586	0.166714

1 Cointegrating Equation(s): Log likelihood -222.2912

Normalized cointegrating coefficients (standard error in parentheses)

EXTERNAL_B...	GDP_GROWTH__ANNUAL__
1.000000	-8.078012
	(2.32213)

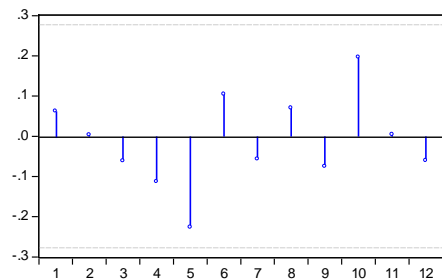
Adjustment coefficients (standard error in parentheses)

D(EXTERNAL_...	0.002773
	(0.02446)
D(GDP_GRO...	0.064553
	(0.01971)

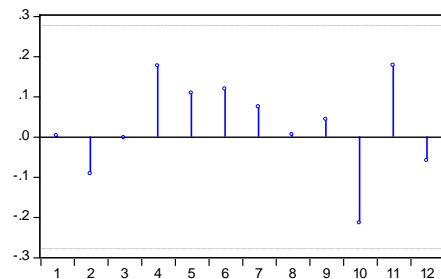
Annex 6.

Autocorrelations with 2 Std.Err. Bounds

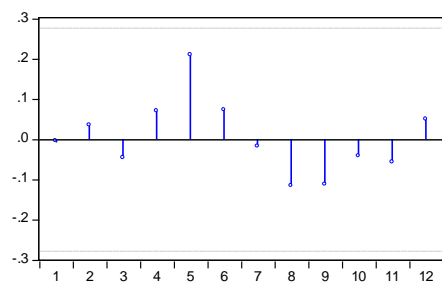
Cor(GDP_GROWTH__ANNUAL____,GDP_GROWTH__ANNUAL____(-i))



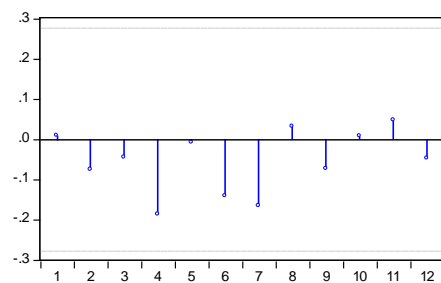
Cor(GDP_GROWTH__ANNUAL____,EXTERNAL_BALANCE_ON_GOOD(i))



Cor(EXTERNAL_BALANCE_ON_GOOD,GDP_GROWTH__ANNUAL____(-i))



Cor(EXTERNAL_BALANCE_ON_GOOD,EXTERNAL_BALANCE_ON_GOOD(i))



Annex 7.

VEC Residual Portmanteau Tests for Autocorrelations

Null Hypothesis: no residual autocorrelations up to lag h

Date: 04/01/18 Time: 21:12

Sample: 1963 2016

Included observations: 52

Lags	Q-Stat	Prob.	Adj Q-Stat	Prob.	df
1	0.223167	NA*	0.227543	NA*	NA*
2	1.197123	0.9770	1.240457	0.9748	6
3	1.798701	0.9977	1.878866	0.9972	10
4	4.940950	0.9866	5.282970	0.9815	14
5	9.335871	0.9514	10.14544	0.9271	18
6	11.98166	0.9578	13.13633	0.9293	22
7	13.80968	0.9754	15.24871	0.9526	26
8	14.65538	0.9915	16.24817	0.9805	30
9	16.38351	0.9953	18.33800	0.9869	34
10	20.13271	0.9923	22.97987	0.9740	38
11	22.30974	0.9946	25.74098	0.9772	42
12	23.02905	0.9982	26.67608	0.9899	46
13	40.60918	0.8257	50.11626	0.4688	50
14	42.04856	0.8815	52.08593	0.5486	54
15	45.72902	0.8786	57.25847	0.5028	58
16	50.18609	0.8590	63.69646	0.4165	62
17	54.35094	0.8465	69.88424	0.3485	66
18	56.17175	0.8847	72.66900	0.3901	70
19	57.50723	0.9216	74.77340	0.4530	74
20	60.13647	0.9335	79.04591	0.4456	78

*The test is valid only for lags larger than the VAR lag order.

df is degrees of freedom for (approximate) chi-square distribution

Annex 8.

Roots of Characteristic Polynomial

Endogenous variables: GDP_GROWTH__ANNUAL_...

Exogenous variables:

Lag specification: 1 1

Date: 04/01/18 Time: 21:08

Root	Modulus
1.000000	1.000000
0.570435	0.570435
-0.300692	0.300692
0.124127	0.124127

VEC specification imposes 1 unit root(s).

Annex 9.

VEC Residual Normality Tests

Orthogonalization: Cholesky (Lutkepohl)

Null Hypothesis: residuals are multivariate normal

Date: 04/01/18 Time: 21:40

Sample: 1963 2016

Included observations: 52

Component	Skewness	Chi-sq	df	Prob.
1	0.251083	0.546372	1	0.4598
2	-0.212071	0.389775	1	0.5324
Joint		0.936146	2	0.6262

Component	Kurtosis	Chi-sq	df	Prob.
1	3.585581	0.742962	1	0.3887
2	3.463175	0.464818	1	0.4954
Joint		1.207780	2	0.5467

Component	Jarque-Bera	df	Prob.
1	1.289334	2	0.5248
2	0.854592	2	0.6523
Joint	2.143926	4	0.7093

Annex 10.

VEC Residual Heteroskedasticity Tests: No Cross Terms (only levels and squares)

Date: 04/01/18 Time: 21:41

Sample: 1963 2016

Included observations: 52

Joint test:		
Chi-sq	df	Prob.
12.36673	18	0.8277

Individual components:					
Dependent	R-squared	F(6,45)	Prob.	Chi-sq(6)	Prob.
res1*res1	0.102318	0.854856	0.5351	5.320557	0.5034
res2*res2	0.063256	0.506456	0.8002	3.289308	0.7718
res2*res1	0.099149	0.825458	0.5563	5.155728	0.5240

VEC Residual Heteroskedasticity Tests: Includes Cross Terms

Date: 04/01/18 Time: 21:41

Sample: 1963 2016

Included observations: 52

Joint test:		
Chi-sq	df	Prob.
23.10345	27	0.6794

Individual components:					
Dependent	R-squared	F(9,42)	Prob.	Chi-sq(9)	Prob.
res1*res1	0.124695	0.664805	0.7354	6.484115	0.6907
res2*res2	0.092346	0.474793	0.8832	4.801987	0.8512
res2*res1	0.204048	1.196334	0.3227	10.61050	0.3034
